

## Reciprocity

$$\hat{G}(\omega, \mathbf{x}, \mathbf{y}) = \hat{G}(\omega, \mathbf{y}, \mathbf{x})$$

## Proof of reciprocity (1/2)

We consider the equations satisfied by the Green's function with the source at  $\mathbf{y}_2$  and with the source at  $\mathbf{y}_1$  ( $\mathbf{y}_2 \neq \mathbf{y}_1$ ):

$$\Delta_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) + \frac{\omega^2}{c^2(\mathbf{x})} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) = -\delta(\mathbf{x} - \mathbf{y}_2)$$

$$\Delta_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) + \frac{\omega^2}{c^2(\mathbf{x})} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) = -\delta(\mathbf{x} - \mathbf{y}_1)$$

We multiply the first equation by  $\hat{G}(\omega, \mathbf{x}, \mathbf{y}_1)$  and subtract the second equation multiplied by  $\hat{G}(\omega, \mathbf{x}, \mathbf{y}_2)$ :

$$\begin{aligned} & \nabla_{\mathbf{x}} \cdot \left[ \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \right] \\ &= \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) \delta(\mathbf{x} - \mathbf{y}_1) - \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \delta(\mathbf{x} - \mathbf{y}_2) \end{aligned}$$

We next integrate over the ball  $B(\mathbf{0}, L)$  which contains both  $\mathbf{y}_1$  and  $\mathbf{y}_2$  and use the divergence theorem:

$$\begin{aligned} & \int_{\partial B(\mathbf{0}, L)} \mathbf{n} \cdot \left[ \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \right] d\sigma(\mathbf{x}) \\ &= \hat{G}(\omega, \mathbf{y}_1, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{y}_2, \mathbf{y}_1) \end{aligned}$$

where  $\mathbf{n}$  is the unit outward normal to the ball  $B(\mathbf{0}, L)$ , which is  $\mathbf{n} = \mathbf{x}/|\mathbf{x}|$ .

## Proof of reciprocity (2/2)

$$\begin{aligned} & \int_{\partial B(\mathbf{0}, L)} \mathbf{n} \cdot \left[ \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \right] d\sigma(\mathbf{x}) \\ &= \hat{G}(\omega, \mathbf{y}_1, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{y}_2, \mathbf{y}_1) \end{aligned}$$

where  $\mathbf{n} = \mathbf{x}/|\mathbf{x}|$ .

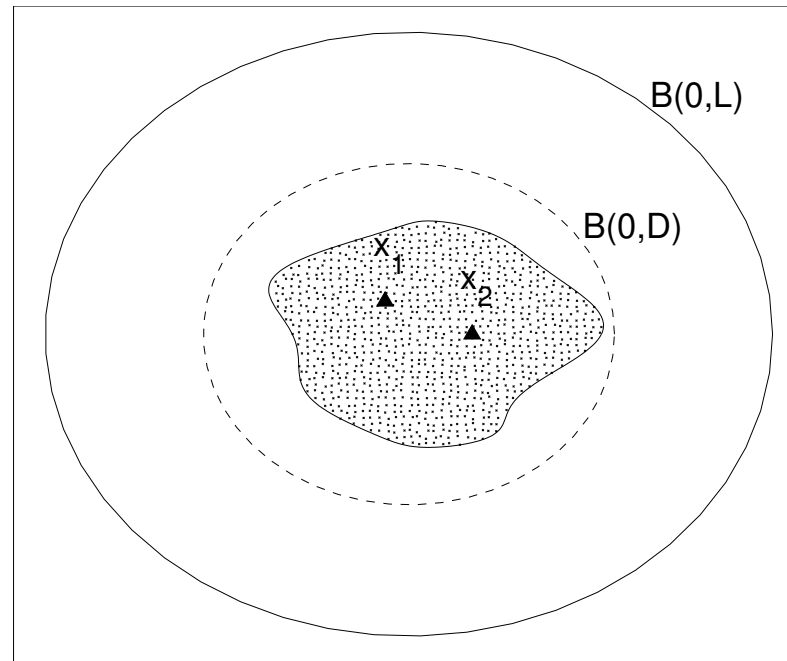
If  $\mathbf{x} \in \partial B(\mathbf{0}, L)$  and  $L \rightarrow \infty$ , then by the Sommerfeld radiation condition:

$$\mathbf{n} \cdot \nabla_{\mathbf{x}} \hat{G}(\omega, \mathbf{x}, \mathbf{y}) = i \frac{\omega}{c_0} \hat{G}(\omega, \mathbf{x}, \mathbf{y}) + o\left(\frac{1}{L}\right)$$

Therefore, for  $L \rightarrow \infty$ ,

$$\begin{aligned} & \hat{G}(\omega, \mathbf{y}_1, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{y}_2, \mathbf{y}_1) \\ &= i \frac{\omega}{c_0} \int_{\partial B(\mathbf{0}, L)} \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) - \hat{G}(\omega, \mathbf{x}, \mathbf{y}_2) \hat{G}(\omega, \mathbf{x}, \mathbf{y}_1) d\sigma(\mathbf{x}) \\ &= 0 \end{aligned}$$

## The Helmholtz-Kirchhoff theorem



If the medium is homogeneous (velocity  $c_o$ ) outside  $B(\mathbf{0}, D)$ , then  $\forall \mathbf{x}_1, \mathbf{x}_2 \in B(\mathbf{0}, D)$  we have for  $L \gg D$ :

$$\hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2) - \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2)} = \frac{2i\omega}{c_o} \int_{\partial B(\mathbf{0}, L)} d\sigma(\mathbf{y}) \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{y})} \hat{G}(\omega, \mathbf{x}_2, \mathbf{y})$$

Proof: second Green's identity and Sommerfeld radiation condition.

Useful for: scattering theory, time reversal experiment, and cross correlation.

## Proof of Helmholtz-Kirchhoff theorem (1/2)

Consider

$$\Delta_{\mathbf{y}} \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) + \frac{\omega^2}{c^2(\mathbf{y})} \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) = -\delta(\mathbf{y} - \mathbf{x}_2)$$

$$\Delta_{\mathbf{y}} \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} + \frac{\omega^2}{c^2(\mathbf{y})} \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} = -\delta(\mathbf{y} - \mathbf{x}_1)$$

Multiply the first equation by  $\overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)}$  and subtract the second equation multiplied by  $\hat{G}(\omega, \mathbf{y}, \mathbf{x}_2)$ :

$$\begin{aligned} & \nabla_{\mathbf{y}} \cdot \left[ \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \nabla_{\mathbf{y}} \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) - \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) \nabla_{\mathbf{y}} \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \right] \\ &= \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) \delta(\mathbf{y} - \mathbf{x}_1) - \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \delta(\mathbf{y} - \mathbf{x}_2) \\ &= \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) \delta(\mathbf{y} - \mathbf{x}_1) - \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{y})} \delta(\mathbf{y} - \mathbf{x}_2) \end{aligned}$$

using the reciprocity property  $\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1) = \hat{G}(\omega, \mathbf{x}_1, \mathbf{y})$ .

Integrate over the ball  $B(\mathbf{0}, L)$  and use the divergence theorem:

$$\begin{aligned} & \int_{\partial B(\mathbf{0}, L)} \mathbf{n} \cdot \left[ \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \nabla_{\mathbf{y}} \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) - \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) \nabla_{\mathbf{y}} \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \right] d\sigma(\mathbf{y}) \\ &= \hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2) - \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2)} \end{aligned}$$

where  $\mathbf{n}$  is the unit outward normal to the ball  $B(\mathbf{0}, L)$ , which is  $\mathbf{n} = \mathbf{y}/|\mathbf{y}|$ .

## Proof of Helmholtz-Kirchhoff theorem (2/2)

$$\begin{aligned} & \int_{\partial B(\mathbf{0}, L)} \mathbf{n} \cdot \left[ \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \nabla_{\mathbf{y}} \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) - \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2) \nabla_{\mathbf{y}} \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)} \right] d\sigma(\mathbf{y}) \\ &= \hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2) - \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2)} \end{aligned}$$

where  $\mathbf{n} = \mathbf{y}/|\mathbf{y}|$ .

This equality can be viewed as an application of the second Green's identity.

The Green's function also satisfies the Sommerfeld radiation condition

$$\lim_{|\mathbf{y}| \rightarrow \infty} |\mathbf{y}| \left( \frac{\mathbf{y}}{|\mathbf{y}|} \cdot \nabla_{\mathbf{y}} - i \frac{\omega}{c_o} \right) \hat{G}(\omega, \mathbf{y}, \mathbf{x}_1) = 0$$

uniformly in all directions  $\mathbf{y}/|\mathbf{y}|$ . Substitute  $i(\omega/c_o)\overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_2)}$  for  $\mathbf{n} \cdot \nabla_{\mathbf{y}} \hat{G}(\omega, \mathbf{y}, \mathbf{x}_2)$  in the surface integral over  $\partial B(\mathbf{0}, L)$ , and  $-i(\omega/c_o)\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)$  for  $\mathbf{n} \cdot \nabla_{\mathbf{y}} \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x}_1)}$ , which gives the desired result:

$$\hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2) - \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{x}_2)} = \frac{2i\omega}{c_o} \int_{\partial B(\mathbf{0}, L)} d\sigma(\mathbf{y}) \overline{\hat{G}(\omega, \mathbf{x}_1, \mathbf{y})} \hat{G}(\omega, \mathbf{x}_2, \mathbf{y})$$

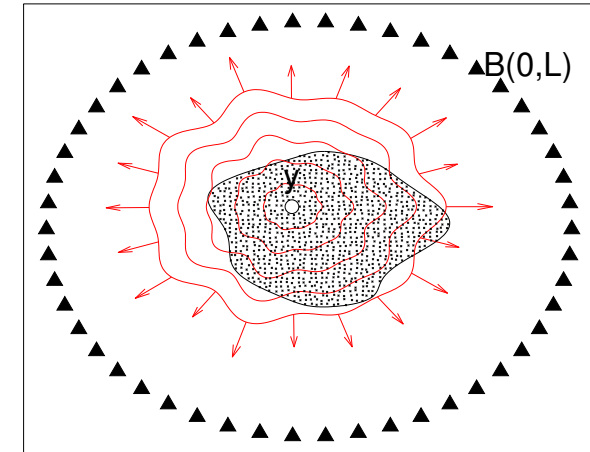
## Time-reversal refocusing on a source (1/2)

First part:

A point source at  $\mathbf{y}$  emits a pulse  $f(t)$ .

The waves are recorded at the surface  $\partial B(\mathbf{0}, L)$ :

$$\hat{u}(\omega, \mathbf{x}) = \hat{G}(\omega, \mathbf{x}, \mathbf{y}) \hat{f}(\omega), \quad \mathbf{x} \in \partial B(\mathbf{0}, L)$$

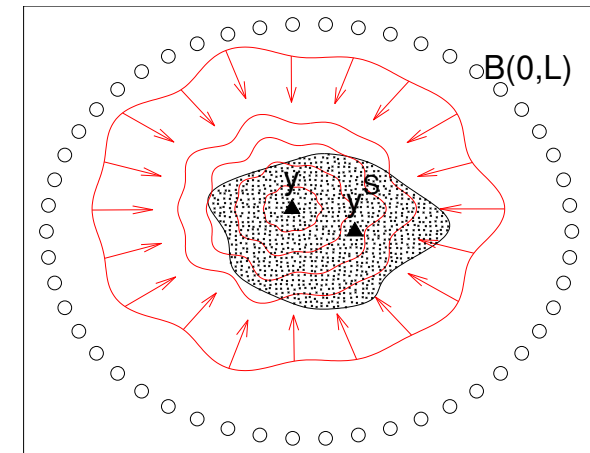


Second part:

The recorded signals are time-reversed and sent back into the medium.

The signal received at  $\mathbf{y}^S$  is

$$\hat{u}_{\text{TR}}(\omega, \mathbf{y}^S) = \int_{\partial B(\mathbf{0}, L)} d\sigma(\mathbf{x}) \hat{G}(\omega, \mathbf{y}^S, \mathbf{x}) \overline{\hat{G}(\omega, \mathbf{x}, \mathbf{y}) \hat{f}(\omega)}$$



## Time-reversal refocusing on a source (2/2)

The signal received at  $\mathbf{y}^S$  is (using reciprocity:  $\hat{G}(\omega, \mathbf{x}, \mathbf{y}) = \hat{G}(\omega, \mathbf{y}, \mathbf{x})$ ):

$$\hat{u}_{\text{TR}}(\omega, \mathbf{y}^S) = \int_{\partial B(\mathbf{0}, L)} d\sigma(\mathbf{x}) \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x})} \hat{G}(\omega, \mathbf{y}^S, \mathbf{x}) \overline{\hat{f}(\omega)}$$

By Helmholtz-Kirchhoff identity:

$$\hat{G}(\omega, \mathbf{y}, \mathbf{y}^S) - \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{y}^S)} = \frac{2i\omega}{c_o} \int_{\partial B(\mathbf{0}, L)} d\sigma(\mathbf{x}) \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{x})} \hat{G}(\omega, \mathbf{y}^S, \mathbf{x})$$

we get

$$\hat{u}_{\text{TR}}(\omega, \mathbf{y}^S) = \frac{\hat{G}(\omega, \mathbf{y}, \mathbf{y}^S) - \overline{\hat{G}(\omega, \mathbf{y}, \mathbf{y}^S)}}{2i\omega/c_o} \overline{\hat{f}(\omega)} = \frac{c_o}{\omega} \text{Im}(\hat{G}(\omega, \mathbf{y}, \mathbf{y}^S)) \overline{\hat{f}(\omega)}$$

[Remember:  $\mathbf{y}$  is the original source location].

In a three-dimensional homogeneous medium:

$$\hat{G}(\omega, \mathbf{y}, \mathbf{y}^S) = \frac{1}{4\pi|\mathbf{y} - \mathbf{y}^S|} e^{i\frac{\omega|\mathbf{y} - \mathbf{y}^S|}{c_o}} \implies \hat{u}_{\text{TR}}(\omega, \mathbf{y}^S) = \frac{1}{4\pi} \text{sinc}\left(\frac{\omega|\mathbf{y} - \mathbf{y}^S|}{c_o}\right) \overline{\hat{f}(\omega)}$$

$\leftrightarrow$  refocusing with a focal spot of diameter  $\lambda/2 = \pi c_o/\omega$  (diffraction limit),

$$\text{sinc}(s) = \frac{\sin(s)}{s}.$$



## Least-square inverse problems

## Inverse problems

- We look for  $\mathbf{m} \in \mathcal{X}$ , the input parameters of a model, given the observed output  $\mathbf{y} \in \mathcal{Y}$ :

$$\mathbf{y} = \mathbf{f}(\mathbf{m}),$$

where  $\mathbf{f} : \mathcal{X} \rightarrow \mathcal{Y}$ ,  $\mathcal{X}$  and  $\mathcal{Y}$  are abstract spaces (Banach).

- Example: calibration of a model.

We have:

- a parametric family of models  $f_{\mathbf{m}} : \mathbb{R}^d \rightarrow \mathbb{R}$ , parameterized by  $\mathbf{m} \in \mathbb{R}^p$ ,
- a sample of observations  $(\mathbf{x}_i, y_i)_{i=1}^n$  with  $y_i = f_{\mathbf{m}^*}(\mathbf{x}_i)(+\text{noise})$ , where  $\mathbf{m}^*$  is unknown.

We want to determine the  $\mathbf{m}$  that best fits the data.

Here  $\mathcal{X} = \mathbb{R}^p$ ,  $\mathcal{Y} = \mathbb{R}^n$ ,  $\mathbf{y} = (y_i)_{i=1}^n$  and  $\mathbf{f}(\mathbf{m}) = (f_{\mathbf{m}}(\mathbf{x}_i))_{i=1}^n$ .

- A problem is ill-posed (in Hadamard's sense) if one of three following events occurs:
  - there is no solution,
  - the solution is not unique,
  - the solution is very sensitive to the data  $\mathbf{y}$ .
- In order to solve an inverse problem, the classical approach is to formulate a least-square minimization problem:

$$\operatorname{argmin}_{\mathbf{m} \in \mathcal{X}} \|\mathbf{y} - \mathbf{f}(\mathbf{m})\|_{\mathcal{Y}}^2.$$

This approach can fail:

- there is no solution in  $\mathcal{X}$ ,
- there are several minima,
- the solution is very sensitive to the data  $\mathbf{y}$ .

## Linear inverse problems

- Here  $\mathcal{X} = \mathbb{R}^p$ ,  $\mathcal{Y} = \mathbb{R}^n$ ,  $\mathbf{f}(\mathbf{m}) = \mathbf{A}\mathbf{m}$  with  $\mathbf{A}$  a  $n \times p$  matrix (with  $n \geq p$ ). We look for

$$\mathcal{S} = \operatorname{argmin}_{\mathbf{m} \in \mathcal{X}} \|\mathbf{y} - \mathbf{A}\mathbf{m}\|^2$$

- Normal equations:

$$\mathbf{m} \in \mathcal{S} \text{ iff } \mathbf{A}^T \mathbf{A}\mathbf{m} = \mathbf{A}^T \mathbf{y}$$

where  $\cdot^T$  stands for transpose (it would be the conjugate transpose in the complex case).

Remark: The normal equations concentrate the problem and the data.

- Use the SVD of  $\mathbf{A} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T$ . The rank  $r$  of  $\mathbf{A}$  is the number of positive singular values. We have

$$\mathcal{S} = \mathbf{A}^+ \mathbf{y} + \operatorname{Ker}(\mathbf{A}) = \left\{ \mathbf{A}^+ \mathbf{y} + \sum_{j=r+1}^p \alpha_j \mathbf{v}_j, \alpha_j \in \mathbb{R}, j = r+1, \dots, p \right\}$$

where  $\mathbf{A}^+$  is the pseudo-inverse of  $\mathbf{A}$ :

$$\mathbf{A}^+ = \mathbf{V}\mathbf{\Sigma}^+ \mathbf{U}^T, \quad \mathbf{\Sigma} = \operatorname{Diag}(\sigma_1^{-1}, \dots, \sigma_r^{-1}, 0, \dots, 0)$$

- Amongst all the solutions (elements of  $\mathcal{S}$ ) we can determine the one with the minimal norm:

$$\mathbf{m}_{\text{LS}} = \underset{\mathbf{m} \in \mathcal{S}}{\operatorname{argmin}} \|\mathbf{m}\|^2 = \mathbf{A}^+ \mathbf{y}$$

If  $\mathbf{y} = \sum_{j=1}^n \beta_j \mathbf{u}_j$  (where the  $\mathbf{u}_j$  are the columns of  $\mathbf{U}$ ), then we have

$$\|\mathbf{y} - \mathbf{A}\mathbf{m}_{\text{LS}}\|^2 = \sum_{j=r+1}^n \beta_j^2$$

- If  $\mathbf{y} = \mathbf{A}\mathbf{m}^{\text{v}}$ ,  $\mathbf{m}^{\text{v}} = \sum_{j=1}^p m_j^{\text{v}} \mathbf{v}_j$  (where the  $\mathbf{v}_j$  are the columns of  $\mathbf{V}$ ), then

$$\mathbf{m}_{\text{LS}} = \sum_{j=1}^r m_j^{\text{v}} \mathbf{v}_j$$

and

$$\|\mathbf{m}_{\text{LS}} - \mathbf{m}^{\text{v}}\|^2 = \sum_{j=r+1}^p (m_j^{\text{v}})^2$$

→ we can determine all coordinates of  $\mathbf{m}^{\text{v}}$  except those which are in  $\operatorname{Ker}(\mathbf{A})$ .

- If  $\mathbf{A}$  has rank  $p$  then  $\mathbf{A}^T \mathbf{A}$  is invertible and  $\mathcal{S}$  is reduced to one point:

$$\hat{\mathbf{m}} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{y}$$

Remark:  $\text{rg}(\mathbf{A}) = p$  iff  $\text{Ker}(\mathbf{A}) = \{\mathbf{0}\}$  iff  $\mathbf{A}^T \mathbf{A}$  is invertible.

- If  $\mathbf{y} = \mathbf{A} \mathbf{m}^v$ , then  $\hat{\mathbf{m}} = \mathbf{m}^v$ . Great !
- If  $\mathbf{y} = \mathbf{A} \mathbf{m}^v + \boldsymbol{\varepsilon}$ , then, using the SVD of  $\mathbf{A} = \mathbf{U} \boldsymbol{\Sigma} \mathbf{V}^T$ , we get

$$\hat{\mathbf{m}} = \mathbf{m}^v + \sum_{j=1}^p \frac{1}{\sigma_j} \varepsilon_j \mathbf{v}_j, \quad \varepsilon_j = \mathbf{v}_j^T \boldsymbol{\varepsilon}.$$

If  $\boldsymbol{\varepsilon} \sim \mathcal{N}(\mathbf{0}, \sigma_{\text{mes}}^2 \mathbf{I})$ , then  $\varepsilon_j$  are i.i.d. with the distribution  $\mathcal{N}(0, \sigma_{\text{mes}}^2)$  and therefore

$$\mathbb{E}[\|\hat{\mathbf{m}} - \mathbf{m}^v\|^2] = \sum_{j=1}^p \frac{\sigma_{\text{mes}}^2}{\sigma_j^2}$$

→ If  $\mathbf{A}$  has a very small singular value, then the error blows up !

## Regularization of ill-posed inverse problems

- A way to regularize the problem is to consider:

$$\operatorname{argmin}_{\mathbf{m} \in \mathcal{X}} \left( \|\mathbf{y} - \mathbf{f}(\mathbf{m})\|_{\mathcal{Y}}^2 + \|\mathbf{m} - \mathbf{m}_0\|_{\mathcal{X}'}^2 \right),$$

with a norm  $\|\cdot\|_{\mathcal{X}'}$  equal to or stronger than the norm  $\|\cdot\|_{\mathcal{X}}$  and  $\mathbf{m}_0 \in \mathcal{X}'$ .

The penalization represents an a priori idea on the structure of the solution.

- Tikhonov regularization:  $\|\cdot\|_{\mathcal{X}'} = \alpha \|\cdot\|_{\mathcal{X}}$  for some  $\alpha > 0$ .

A priori idea: the solution should not have a very large norm  $\|\cdot\|_{\mathcal{X}}$ .

If  $\alpha$  is large enough, then the function to be minimized may become convex.

However: If  $\alpha$  is too large, then the regularized problem becomes significantly different from the original problem.

- $L^2$  regularization (ridge regression).

Here  $\mathcal{X} = \mathbb{R}^p$ ,  $\mathcal{Y} = \mathbb{R}^n$ ,  $\mathbf{f}(\mathbf{m}) = \mathbf{A}\mathbf{m}$  with  $\mathbf{A}$   $n \times p$ -matrix with  $p \leq n$ .

Let  $\alpha > 0$ . We look for

$$\operatorname{argmin}_{\mathbf{m} \in \mathcal{X}} \left( \|\mathbf{y} - \mathbf{A}\mathbf{m}\|^2 + \alpha \|\mathbf{m}\|^2 \right)$$

- Normal equations:

$$\mathbf{m} \in \mathcal{S}_\alpha \text{ iff } (\mathbf{A}^T \mathbf{A} + \alpha \mathbf{I})\mathbf{m} = \mathbf{A}^T \mathbf{y}.$$

- There is a unique solution because  $\mathbf{A}^T \mathbf{A} + \alpha \mathbf{I}$  is positive-definite:

$$\hat{\mathbf{m}}_\alpha = (\mathbf{A}^T \mathbf{A} + \alpha \mathbf{I})^{-1} \mathbf{A}^T \mathbf{y}$$

- By using the SVD of  $\mathbf{A} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T$ , we find

$$\hat{\mathbf{m}}_\alpha = \sum_{j=1}^p \frac{\sigma_j}{\sigma_j^2 + \alpha} y_j \mathbf{v}_j, \quad y_j = \mathbf{u}_j^T \mathbf{y}$$



$$\hat{\mathbf{m}}_\alpha = \sum_{j=1}^p \frac{\sigma_j}{\sigma_j^2 + \alpha} y_j \mathbf{v}_j, \quad y_j = \mathbf{u}_j^T \mathbf{y}$$

- If  $\mathbf{y} = \mathbf{A}\mathbf{m}^\vee$ , with  $\mathbf{m}^\vee = \sum_{j=1}^p m_j^\vee \mathbf{v}_j$ ,  $m_j^\vee = \mathbf{v}_j^T \mathbf{m}$ , then  $\mathbf{y} = \sum_{j=1}^n y_j \mathbf{u}_j$  with  $y_j = \sigma_j m_j^\vee$ . Therefore

$$\hat{\mathbf{m}}_\alpha = \sum_{j=1}^p \frac{\sigma_j^2}{\sigma_j^2 + \alpha} m_j^\vee \mathbf{v}_j$$

and the error is

$$\hat{\mathbf{m}}_\alpha - \mathbf{m}^\vee = - \sum_{j=1}^p \frac{\alpha}{\sigma_j^2 + \alpha} m_j^\vee \mathbf{v}_j$$

$$\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^\vee\|^2 = \sum_{j=1}^p \frac{\alpha^2}{(\sigma_j^2 + \alpha)^2} (m_j^\vee)^2$$

→ It seems that we should take  $\alpha = 0$  !

$$\hat{\mathbf{m}}_\alpha = \sum_{j=1}^p \frac{\sigma_j}{\sigma_j^2 + \alpha} y_j \mathbf{v}_j, \quad y_j = \mathbf{u}_j^T \mathbf{y}$$

- If  $\mathbf{y} = \mathbf{A}\mathbf{m}^v + \boldsymbol{\varepsilon}$ , with  $\boldsymbol{\varepsilon} \sim \mathcal{N}(\mathbf{0}, \sigma_{\text{mes}}^2 \mathbf{I})$ , then we find

$$\hat{\mathbf{m}}_\alpha = \sum_{j=1}^p \frac{\sigma_j^2}{\sigma_j^2 + \alpha} m_j^v \mathbf{v}_j + \sum_{j=1}^p \frac{\sigma_j}{\sigma_j^2 + \alpha} \varepsilon_j \mathbf{v}_j, \quad \varepsilon_j = \mathbf{u}_j^T \boldsymbol{\varepsilon}$$

Since  $\varepsilon_j$  are i.i.d. with the distribution  $\mathcal{N}(0, \sigma_{\text{mes}}^2)$ ,

$$\begin{aligned} \mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^v\|^2] &= \mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbb{E}[\hat{\mathbf{m}}_\alpha]\|^2] + \|\mathbb{E}[\hat{\mathbf{m}}_\alpha] - \mathbf{m}^v\|^2 \\ &= \underbrace{\sum_{j=1}^p \frac{\sigma_j^2}{(\sigma_j^2 + \alpha)^2} \sigma_{\text{mes}}^2}_{\text{variance}} + \underbrace{\sum_{j=1}^p \frac{\alpha^2}{(\sigma_j^2 + \alpha)^2} (m_j^v)^2}_{\text{bias}} \end{aligned}$$

→ The variance term does not blow up when there are small singular values, it is bounded by  $p\sigma_{\text{mes}}^2/(4\alpha)$  uniformly w.r.t.  $(\sigma_j)_{j=1}^p$ .

→ By allowing for a small bias, one can strongly reduce the variance.

↔ We look for adjusting the parameter  $\alpha$  in order to minimize  $\mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^v\|^2]$  (bias-variance trade-off).

$$\mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^\vee\|^2] = \sum_{j=1}^p \frac{\sigma_j^2}{(\sigma_j^2 + \alpha)^2} \sigma_{\text{mes}}^2 + \sum_{j=1}^p \frac{\alpha^2}{(\sigma_j^2 + \alpha)^2} (m_j^\vee)^2$$

- The optimal  $\alpha$  depends on  $\mathbf{m}^\vee$  and  $\mathbf{A}$ , it exists and it is positive.

→ one must regularize !

*Proof (of positivity).* The function  $\alpha \mapsto \mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^\vee\|^2]$  is continuous, decaying close to 0,

$$\mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^\vee\|^2] = \left[ \sum_{j=1}^p \sigma_j^{-2} \sigma_{\text{mes}}^2 \right] - \alpha \left[ 2 \sum_{j=1}^p \sigma_j^{-4} \sigma_{\text{mes}}^2 \right] + O(\alpha^2), \quad \alpha \rightarrow 0^+,$$

and increasing at infinity:

$$\mathbb{E}[\|\hat{\mathbf{m}}_\alpha - \mathbf{m}^\vee\|^2] = \left[ \sum_{j=1}^p (m_j^\vee)^2 \right] - \alpha^{-1} \left[ 2 \sum_{j=1}^p \sigma_j^2 (m_j^\vee)^2 \right] + O(\alpha^{-2}), \quad \alpha \rightarrow +\infty.$$

- In practice (Morozov rule): choose  $\alpha$  so that  $\|\mathbf{y} - \mathbf{A}\hat{\mathbf{m}}_\alpha\|^2 \simeq \mathbb{E}[\|\boldsymbol{\varepsilon}\|^2] = n\sigma_{\text{mes}}^2$ .

→ do not try to adjust the model with an accuracy higher than the noise level (overfitting).

- $L^0$  regularization:

$$\operatorname{argmin}_{\mathbf{m} \in \mathcal{X}} \left( \|\mathbf{y} - \mathbf{A}\mathbf{m}\|^2 + \alpha \|\mathbf{m}\|_0 \right),$$

with  $\|\mathbf{m}\|_0 = \sum_{j=1}^p \mathbf{1}_{m_j \neq 0} = \operatorname{Card}\{j = 1, \dots, p, m_j \neq 0\}$ .

- We look for a *sparse* solution (parameter selection).
- The problem is numerically very challenging.

- $L^1$  regularization (lasso regression = Least Absolute Shrinkage and Selection Operator):

$$\operatorname{argmin}_{\mathbf{m} \in \mathcal{X}} \left( \|\mathbf{y} - \mathbf{A}\mathbf{m}\|^2 + \alpha \|\mathbf{m}\|_1 \right),$$

with  $\|\mathbf{m}\|_1 = \sum_{j=1}^p |m_j|$ .

- For “good” matrices  $\mathbf{A}$  (satisfying the RIP - Restricted Isometry Property), the solution of the  $L^1$ -regularized problem is the same as the solution of the  $L^0$ -regularized problem.
- The problem is numerically challenging (not differentiable function), but not impossible (in particular, when the solution is sparse).

## Bayesian approach of inverse problems

- Principle: We do not look for a unique answer, but a probability measure on  $\mathcal{X}$  which gives the likelihood of the states  $\mathbf{m}$  given  $\mathbf{y}$ .
- We assume that:
  - $\mathcal{X} = \mathbb{R}^p$  and  $\mathcal{Y} = \mathbb{R}^n$ .
  - We have a priori information on the most likely states  $\mathbf{m}$  in the form of an a priori distribution on  $\mathcal{X}$  with density  $\pi_0$  (w.r.t. the Lebesgue measure on  $\mathcal{X}$ ).
  - The observations are noisy:

$$\mathbf{y} = \mathbf{f}(\mathbf{m}) + \boldsymbol{\eta},$$

where  $\boldsymbol{\eta}$  is a random variable taking values in  $\mathcal{Y}$  with density  $\rho$  (w.r.t. the Lebesgue measure on  $\mathcal{Y}$ ).

→ The likelihood (distribution of  $\mathbf{y}$  given  $\mathbf{m}$ ) has density  $\rho(\mathbf{y} - \mathbf{f}(\mathbf{m}))$ .

→ Bayes theorem: the a posteriori distribution of  $\mathbf{m}$  given  $\mathbf{y}$  has density:

$$\pi_{\mathbf{y}}(\mathbf{m}) = \frac{\rho(\mathbf{y} - \mathbf{f}(\mathbf{m}))\pi_0(\mathbf{m})}{\int_{\mathbb{R}^p} \rho(\mathbf{y} - \mathbf{f}(\mathbf{m}'))\pi_0(\mathbf{m}')d\mathbf{m}'}$$

## Bayesian approach - Gaussian case

$$\mathbf{y} = \mathbf{f}(\mathbf{m}) + \boldsymbol{\eta}$$

- We assume, moreover, that the a priori distribution of  $\mathbf{m}$  is  $\mathcal{N}(\mathbf{m}_0, \boldsymbol{\Sigma}_0)$  and that the distribution of  $\boldsymbol{\eta}$  is  $\mathcal{N}(\mathbf{0}, \boldsymbol{\Gamma})$ , with invertible real matrices  $\boldsymbol{\Sigma}_0$  and  $\boldsymbol{\Gamma}$ .

$$\pi^0(\mathbf{m}) = \frac{1}{(2\pi)^{p/2} \text{Det}(\boldsymbol{\Sigma}_0)^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{m} - \mathbf{m}_0)^T \boldsymbol{\Sigma}_0^{-1}(\mathbf{m} - \mathbf{m}_0)\right)$$

- The a posteriori distribution of  $\mathbf{m}$  given  $\mathbf{y}$  has density:

$$\pi_{\mathbf{y}}(\mathbf{m}) \approx \exp\left(-\frac{1}{2}\|\mathbf{y} - \mathbf{f}(\mathbf{m})\|_{\boldsymbol{\Gamma}}^2 - \frac{1}{2}\|\mathbf{m} - \mathbf{m}_0\|_{\boldsymbol{\Sigma}_0}^2\right),$$

where

$$\|\mathbf{y}\|_{\boldsymbol{\Gamma}}^2 = \mathbf{y}^T \boldsymbol{\Gamma}^{-1} \mathbf{y} = (\boldsymbol{\Gamma}^{-1/2} \mathbf{y})^T \boldsymbol{\Gamma}^{-1/2} \mathbf{y} = \|\boldsymbol{\Gamma}^{-1/2} \mathbf{y}\|^2.$$

- The Maximum A Posteriori (MAP) (mode of the a posteriori distribution of  $\mathbf{m}$  given  $\mathbf{y}$ ) is:

$$\operatorname{argmin}_{\mathbf{m} \in \mathbb{R}^p} \left( \|\mathbf{y} - \mathbf{f}(\mathbf{m})\|_{\boldsymbol{\Gamma}}^2 + \|\mathbf{m} - \mathbf{m}_0\|_{\boldsymbol{\Sigma}_0}^2 \right).$$

↔ We recover the  $L^2$ -regularized minimization problem.

## Bayesian approach - Laplace case

$$\mathbf{y} = \mathbf{f}(\mathbf{m}) + \boldsymbol{\eta}$$

- We assume, moreover, that the a priori distribution of  $\mathbf{m}$  is a Laplace distribution with density

$$\pi^0(\mathbf{m}) = \lambda^p \exp\left(-\frac{\lambda}{2}\|\mathbf{m}\|_1\right)$$

with  $\lambda > 0$ , and that the distribution of  $\boldsymbol{\eta}$  is  $\mathcal{N}(\mathbf{0}, \boldsymbol{\Gamma})$ .

- The a posteriori distribution of  $\mathbf{m}$  given  $\mathbf{y}$  has density

$$\pi_{\mathbf{y}}(\mathbf{m}) \approx \exp\left(-\frac{1}{2}\|\mathbf{y} - \mathbf{f}(\mathbf{m})\|_{\boldsymbol{\Gamma}}^2 - \frac{\lambda}{2}\|\mathbf{m}\|_1\right).$$

- The Maximum A Posteriori (MAP) is:

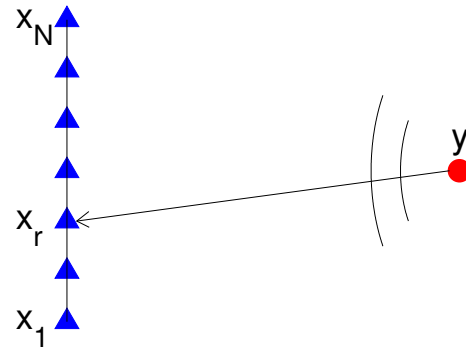
$$\operatorname{argmin}_{\mathbf{m} \in \mathbb{R}^p} \left( \|\mathbf{y} - \mathbf{f}(\mathbf{m})\|_{\boldsymbol{\Gamma}}^2 + \lambda \|\mathbf{m}\|_1 \right).$$

→ We recover the  $L^1$ -regularized minimization problem. We understand (from the cusp of the Laplace density) that the MAP favors the solutions with entries equal to 0.

# Source imaging



## Source imaging: data acquisition



Passive array  $\iff$  The sensors are receivers.

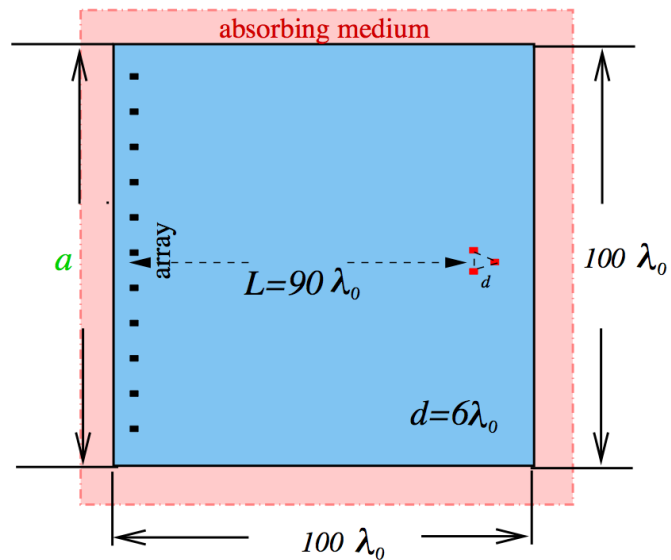
The source  $\mathbf{y}$  emits a short pulse.

The sensors  $(\mathbf{x}_r)_{r=1,\dots,N}$  record the waves.

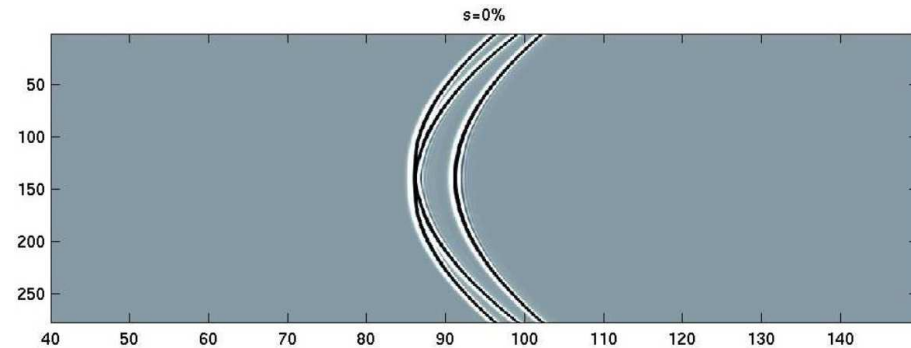
The data set is  $(u(t, \mathbf{x}_r))_{t \in \mathbb{R}, r=1,\dots,N}$ .

Goal: find the source position  $\mathbf{y}$  (more generally, find the *source* region).

## Source imaging: simulation

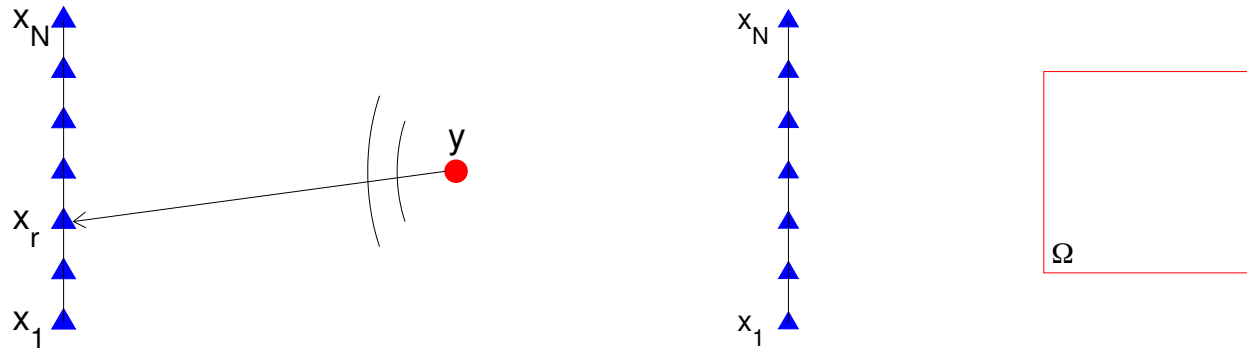


Configuration



Data set  $(u(t, \mathbf{x}_r))_{80 \leq t \leq 200, r=1, \dots, 270}$

## Source imaging: imaging function



Data acquisition

Search region

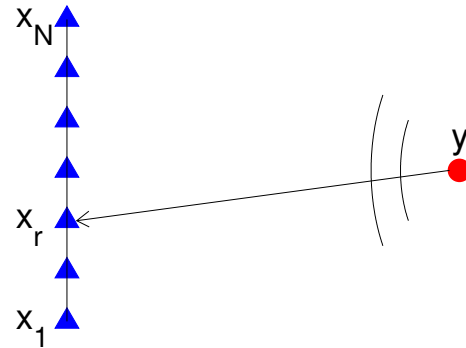
Goal: find the source point  $\mathbf{y}$  (more generally, find the *source* region).

The data set is  $(u(t, \mathbf{x}_r))_{t \in \mathbb{R}, r=1, \dots, N}$ .

$\Leftrightarrow$  Given the data set, build an imaging function in the search region  $\Omega \subset \mathbb{R}^3$ :

$$\mathcal{I} : \begin{cases} \Omega \rightarrow \mathbb{R}^+ \\ \mathbf{y}^S \mapsto \mathcal{I}(\mathbf{y}^S) \end{cases} \quad \text{which plots an image of the search region.}$$

## Source imaging - the linear forward operator



The source term is of the form  $n(t, \mathbf{y}) = \rho_{\text{real}}(\mathbf{y})\delta(t)$ .

Goal: find the source function  $\rho_{\text{real}}$ .

Here: the Green's function is known.

The data set is  $\hat{\mathbf{u}} = (\hat{u}(\omega, \mathbf{x}_r))_{\omega \in \mathcal{B}, r=1, \dots, N}$  with  $\hat{u}(\omega, \mathbf{x}_r) = \int_{\Omega} \hat{G}(\omega, \mathbf{x}_r, \mathbf{y}) \rho_{\text{real}}(\mathbf{y}) d\mathbf{y}$ .

We define

$$[\hat{\mathbf{A}}\rho](\omega, \mathbf{x}_r) = \int_{\Omega} \hat{G}(\omega, \mathbf{x}_r, \mathbf{y}) \rho(\mathbf{y}) d\mathbf{y}.$$

$\hat{\mathbf{A}}$  is the linear operator that maps the source function to the array data  $\hat{\mathbf{u}}$ :

$$\hat{\mathbf{u}} = \hat{\mathbf{A}}\rho_{\text{real}}$$